# 3.2 Likelihood functions:

### 1. Poisson distribution

The distribution of a Poisson random variable Y is

$$P(Y = y) = \frac{exp(-\mu)\mu^{y}}{v!}, y = 0, 1, 2, \cdots$$

The cumulant generating function is

$$\kappa_{\rm Y} = \mu[exp(t) - 1].$$

Thus,

$$E(Y) = \kappa_T'(0) = [\mu \cdot exp(t)]_{t=0} = \mu$$

and

$$E(Y) = \kappa_T''(0) = [\mu \cdot exp(t)]_{t=0} = \mu.$$

## 2. The Poisson log-likelihood function

The Poisson log-likelihood function for independent  $Y_i \sim P(\mu_i)$ ,  $i = 1, \dots, n$ , is

$$l(\mu_1, \mu_2, \dots, \mu_n) \propto \sum_{i=1}^n [y_i \cdot log(\mu_i) - \mu_i],$$

where  $E(Y_i) = \mu_i$ . The deviance function is

$$\begin{aligned} &D(y_1, y_2, \cdots, y_n | \widetilde{\mu}_1, \widetilde{\mu}_2, \cdots, \widetilde{\mu}_n) \\ &= 2 \sum_{i=1}^n \left[ y_i log\left(\frac{y_i}{\widetilde{\mu}_i}\right) - (y_i - \widetilde{\mu}_i) \right] \\ &= 2 \sum_{i=1}^n \left[ y_i log\left(\frac{y_i}{\widetilde{\mu}_i}\right) \right] - 2 \sum_{i=1}^n (y_i - \widetilde{\mu}_i) \end{aligned}$$

where  $\ \widetilde{\mu}_i = g^{-1} ig( x_i \widehat{oldsymbol{eta}} ig)$  is the estimate of  $\ E(Y_i) = \mu_i.$ 

### Note:

If a constant term (the intercept) is included in the model, it can be shown that

$$\sum_{i=1}^n (y_i - \widetilde{\mu}_i) = 0.$$

Thus, the deviance function can be reduced to

$$D(y_1, y_2, \cdots, y_n | \widetilde{\mu}_1, \widetilde{\mu}_2, \cdots, \widetilde{\mu}_n) = 2 \sum_{i=1}^n \left[ y_i log\left(\frac{y_i}{\widetilde{\mu}_i}\right) \right].$$

Note:

The deviance function is closely related to Pearson's statistic. Since

$$\frac{y_i - \widetilde{\mu}_i}{\widetilde{\mu}_i} = \varepsilon_i \Longleftrightarrow \frac{y_i}{\widetilde{\mu}_i} = 1 + \varepsilon_i \Longleftrightarrow y_i = \widetilde{\mu}_i (1 + \varepsilon_i),$$

then

$$\begin{split} y_i \cdot log\left(\frac{y_i}{\widetilde{\mu}_i}\right) - (y_i - \widetilde{\mu}_i) &= \widetilde{\mu}_i(1 + \varepsilon_i) \cdot log\left(\frac{y_i}{\widetilde{\mu}_i}\right) - \widetilde{\mu}_i\varepsilon_i \\ &= \widetilde{\mu}_i \big[ (1 + \varepsilon_i) \cdot log(1 + \varepsilon_i) - \varepsilon_i \big] \\ &= \widetilde{\mu}_i \left[ (1 + \varepsilon_i) \left( \varepsilon_i - \frac{\varepsilon_i^2}{2} + \cdots \right) - \varepsilon_i \right] \\ &= \widetilde{\mu}_i \left[ \left( \varepsilon_i - \frac{\varepsilon_i^2}{2} + \varepsilon_i^2 - \frac{\varepsilon_i^3}{2} + \cdots \right) - \varepsilon_i \right] \\ &= \widetilde{\mu}_i \left( \frac{\varepsilon_i^2}{2} + \cdots \right) \\ &\approx \frac{\widetilde{\mu}_i \varepsilon_i^2}{2}, \end{split}$$

where

$$log(1+\varepsilon_i)=\varepsilon_i-\frac{\varepsilon_i^2}{2}+\cdots.$$

Thus, as the model includes the intercept,

$$\begin{split} D(y_1, y_2, \cdots, y_n | \widetilde{\mu}_1, \widetilde{\mu}_2, \cdots, \widetilde{\mu}_n) &= 2 \sum_{i=1}^n \left[ y_i log\left(\frac{y_i}{\widetilde{\mu}_i}\right) \right] \\ &\approx \sum_{i=1}^n \frac{(y_i - \widetilde{\mu}_i)^2}{\widetilde{\mu}_i} \end{split}$$

**≡** Pearson's statistic

#### Note:

 $\sigma^2$  can be estimated by

$$\widetilde{\sigma}^2 = \frac{1}{n-p} \sum_{i=1}^n \frac{(y_i - \widetilde{\mu}_i)^2}{\widetilde{\mu}_i} = \frac{X^2}{n-p'}$$

 $X^2$  is the sum of Pearson's residuals, i.e., Pearson's statistic.